LIMIT THEOREMS OF PROBABILITY THEORY

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Summary

Some basic theory of sums of random variables with increasing number of terms is presented. Different types of convergence are treated.

The laws of large numbers, the law of iterated logarithm, the central limit theorem and the classical summation theory are given, mainly for sums of independent random variables, and also refinements on these theorems.

Local limit theorems, asymptotic expansions, large deviations results and limit distributions of normalized extremes and order statistics are considered, too.

1. Introduction and Preliminaries

Probability theory is motivated by the idea, that the unknown probability p of an event A is approximately equal to r/n, if n trials result in r realisation of the event A, and the approximation improves with increasing n. Limit theorems in probability theory and statistics are regarded as results giving convergence of sequences of random variables or their distribution functions. Since sequences of random variables are sequences of functions with random influences, different modes of convergence are involved. The law of large numbers and the central limit theorem are the most important limit theorems. They are parts of the classical summation theory, investigating the possible limit distributions for the distributions of certain sums of random variables.

1.1. Sequences of Events and Their Probabilities

Let (Ω, \mathbf{A}, P) be a *probability space*, where Ω is a set of elements ω , \mathbf{A} is a σ -algebra of subsets (here called events) of the set Ω , and P is a probability measure defined on \mathbf{A} . Let $\{A_n\}_{n\geq 1} \in \mathbf{A}$ be a sequence of events. (See "Mathematical Foundations and Interpretations of Probability").

Proposition 1.1. Boole's inequality: For events $\{A_n\}_{n\geq 1} \in A$,

$$P(\bigcup_{k=1}^{\infty} A_k) \leq \sum_{k=1}^{\infty} P(A_k) .$$
⁽¹⁾

Since P is countable additive, the equal sign in (1) holds for pairwise disjoint events, i.e. if $A_i \cap A_j = \emptyset$ for $i \neq j$. Define the following events belonging to **A**:

$$\limsup_{n \to \infty} A_n = \bigcap_{k=1}^{\infty} \bigcup_{n=k}^{\infty} A_n = \{A_n, \text{ i.o. }\} \text{ and } \liminf_{n \to \infty} A_n = \bigcup_{k=1}^{\infty} \bigcap_{n=k}^{\infty} A_n$$
(2)

The set $\limsup_n A_n$, denoted by $\{A_n, i.o.\}$ is the set of events ω such that $\omega \in A_n$ for *infinitely many* values of n, where i.o. abbreviates *"infinitely often"*. The set $\liminf_n A_n$ is the set of such events ω , that $\omega \in A_n$ for all but finitely many values of n.

Proposition 1.2. For events $\{A_n\}_{n \ge 1} \in \mathbf{A}$ the following hold:

 $P(\liminf_{n \to \infty} A_{n}) \le \liminf_{n \to \infty} P(A_{n}) \text{ and}$ $\limsup_{n \to \infty} P(A_{n}) \le P(\limsup_{n \to \infty} A_{n}). \tag{3}$

The first inequality in (3) is a consequence of Fatou's lemma for probabilities.

Proposition 1.3. *Borel-Cantelli lemma:* Suppose $\{A_n\}_{n \ge 1} \in \mathbf{A}$.

a) If $\sum_{n=1}^{\infty} P(\mathbf{A}_n) < \infty$, then $P(\mathbf{A}_n, \text{i.o.}) = 0$.

b) Let A₁, A₂,... be independent events such that $\sum_{n=1}^{\infty} P(A_n) = \infty$. Then $P(A_n, i.o.) = 1$.

Proposition 1.4. *Borel zero-one criterion:* If the events in the sequence $\{A_n\}_{n \ge 1} \in \mathbf{A}$ are independent, then $P(A_n, i.o.) = 0$ or 1 according as $\sum_{n=1}^{\infty} P(A_n) < \infty$ or $= \infty$.

1.2. Inequalities for Sums of Random Variables

Let $X, X_1, X_2,...$ be random variables on a common probability space (Ω, \mathbf{A}, P) . Denote the *mathematical expectation*, the *variance* and the *d-th order absolute moment* of X by EX, Var X and $E|X|^d$, respectively, if they exist. The space $L^d = L^d(\Omega, \mathbf{A}, P), \ 0 < d < \infty$, denotes the set of random variables X such that $E|X|^d < \infty$. The usual metric in the space L^d is given by $d(X,Y) = ||X - Y||_d$ with $||X||_d = E|X|^d$ or $(E|X|^d)^{1/d}$ according as 0 < d < 1 or $d \ge 1$. There is a type of very important inequalities which are collected by the *Markov-inequality:*

 $P(|X| \ge \varepsilon) \le (g(\varepsilon))^{-1} E g(X) \text{ for any even non-degreasing function } g \ge 0 \text{ and every}$ $\varepsilon > 0.$ (4)

With $g(x) = x^2$ the Markov inequality implies the *Bienaymé-Chebyshev inequality*:

$$P(|X - EX| \ge \varepsilon) \le \varepsilon^{-2} \text{ Var } X \quad \text{for every } \varepsilon > 0, \qquad (5)$$

estimating the probability of deviation of a random variable from its expectation by its variance.

Consider the partial sum $S_n = X_1 + \dots + X_n$ from the sequence $\{X_n\}_{n \ge 1}$.

Example 1.1. Weak law of large numbers: Let X_1, X_2, \dots, X_n be independent and identically distributed (Two or more random variables are identically distributed, if they have the same distribution.) with finite variance $0 < \sigma^2 = \operatorname{Var} X_1 < \infty$. Then $E S_n = n \mu$, $\operatorname{Var} S_n = n \sigma^2$ and by (5)

$$P(|S_n/n - EX_1| \ge \varepsilon) \le \sigma^2 \varepsilon^{-2} n^{-1} \text{ for any } \varepsilon > 0.$$
(6)

Hence, the probability of the event, that the arithmetic mean $\overline{X_n} = S_n / n$ differs from the expectation of the summands EX_1 by more than \mathcal{E} , tends to zero.

Let $\{X_n\}_{n\geq 1} \in L^2$ be a sequence of independent random variables with $EX_k = 0$ and $\sigma_k^2 = \operatorname{Var} X_k < \infty$, $k = 1, \dots, n$. An useful tool in probability theory is the *Hàjek-Rényi* inequality: Suppose $0 < c_n \le c_{n-1} \le \dots \le c_1$. Then, for all x > 0 and every integer 0 < m < n,

$$P(\max_{m \le k \le n} |S_k| \ge x) \le x^{-2} (c_m^2 \sum_{k=1}^m \sigma_k^2 + \sum_{k=m+1}^n c_k^2 \sigma_k^2).$$
(7)

In case $c_1 = \cdots = c_n = 1$ one find the *Kolmogorov inequality*:

$$P(\max_{1 \le k \le n} | S_k | \ge x) \le x^{-2} \sum_{k=1}^n \sigma_k^2$$

Consider the Bernstein condition: There exists a positive constant H such that

$$|EX_k^m| \le \frac{m!}{2}\sigma_k^2 H^{m-2} \text{ for all integers } m \ge 2 \text{ and } k = 1, 2..., n, \qquad (9)$$

bounding the growth of the moments of X_k . Bernstein's condition (9) implies exponential estimates for the partial sum S_n , the *Bernstein inequalities*: Put $b_n^2 = \sigma_1^2 + \dots + \sigma_n^2$, then

(8)

$$\max\{P(S_n \le -x), P(S_n \ge x)\} \le \begin{cases} \exp\{-x^2/(4b_n^2)\} & \text{if } 0 \le x \le b_n^2/H, \\ \exp\{-x/(4H)\} & \text{if } x \ge b_n^2/H. \end{cases}$$
(10)

The Bernstein inequalities are rather powerful, leading to exponentially fast convergence rates as shown in the following. If the random variables $X_1,...,X_n$ with zero mean are uniformly bounded, i.e. if there is a constant *C* such that $P(|X_k| \le C) = 1$ for k = 1,...,n, then Bernstein's condition (9) is satisfied with H = C. In case of identically distributed random variables $X_1,...,X_n$, the Bernstein condition (9) is a consequence of the *Cramér condition*: There exists a positive constant *a* such that

$$E\exp\{a \mid X \mid\} < \infty, \tag{11}$$

ensuring the existence of the *moment generating function* $M(h) = E \exp\{hX\}$ for $|h| \le a$, which leads to the existence of all order moments of the underlying random variable *X*.

Example 1.2. Large deviation estimate for the arithmetic mean: Consider a sequence $\{X_n\}_{n\geq 1}$ of independent and identically distributed random variables with $P(X_1 = 1) = p = 1 - P(X_1 = 0)$. Then the partial sum $S_n = X_1 + \dots + X_n$ is binomial distributed with success probability p, 0 . In order to apply Bernstein'sinequalities (10) the conditions (9) have to be proved for the random variable $(X_1 - p)$, which has zero mean, as variable now well as the random $(X_1 - p) + \dots + (X_n - p) = S_n - np$. Since $|X_1 - p| \le 1$, Bernstein's condition (9) holds with H = 1. Let $\overline{X_n} = S_n / n$ denote the arithmetic mean of the first *n* random variables from the given sequence. Using $4p(1-p) \le 1$, Bernstein inequalities (10) with $x = \varepsilon n$ for some $0 < \varepsilon < p(1 - p)$ imply an exponential bound for the deviation of sample mean $\overline{X_n}$ from success probability *p*:

$$\max\{P(\overline{X_n} - p \le -\varepsilon), \ P(\overline{X_n} - p \ge \varepsilon)\} \le \exp\{-n\varepsilon^2\}.$$
(12)

Hence, $P(|\overline{X_n} - p| \ge \varepsilon)$ tends to zero exponentially fast as $n \to \infty$. Inequalities like (12) are known as *large deviation estimations* in the law of large numbers.

Let $X_k \in L^r$, i.e. $E |X_k|^r < \infty$ for some $r \ge 2$. Define $M_{r,n} = \sum_{k=1}^n E |X_k|^r$, then by the *Fuk-Nagaev inequality*:

$$P(S_n \ge x) \le (1+2/r)^r M_{r,n} x^{-r} + \exp\{-2(r-2)e^{-r}x^2 B_n^{-2}\} \text{ for } x > 0.$$
(13)

Finally, moments $E |S_n|^r$ may be estimated by the c_r -inequality

$$E |S_n|^r \le c_r M_{r,n}$$
 with $c_r = n^{r-1}$ if $r \ge 1$ or $c_r = 1$ if $r \le 1$. (14)

For independent sequences the Rosenthal inequalities give upper and lower bounds:

$$2^{-r} \max\{M_{r,n}, B_n^r\} \le E(|S_n|^r) \le 2^{r \cdot r} \max\{M_{r,n}, B_n^r\}$$
(15)

with $r \ge 2$ in the first and r > 1 in the second of the inequalities.

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Biographical Sketch

Gerd Christoph is professor of probability theory at the Department of Mathematics at the Otto-von-Guericke University of Magdeburg/Germany. He was born October 30, 1948 in Jena/Germany and is married with 3 children. He studied mathematics at the State University in Leningrad (now St. Petersburg)/Russia from 1967 to 1972 and worked as research assistant at the Technical University of Dresden from 1973 to 1987, where he received the degrees Dr. rer. nat. in 1978 and Dr. rer. nat. habil. in 1987. From 1987 to 1989 he was a lecturer at the Department of Mathematics at the Addis Ababa University in Ethiopia, from 1989 to 1990 a docent and since 1990 he has been professor in Magdeburg. Fields of his interest are limit theorems in probability theory, stable and discrete stable limit distributions, characterization, and stability problems for quadratic forms, extreme value theorems, and asymptotic methods in statistics.